

## A BAYESIAN STATISTICAL ANALYSIS OF THE DNA CONTAMINATION SCENARIO

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**ABSTRACT:** When an unknown DNA profile recovered from crime-scene evidence is later found to match someone's known DNA profile, a question that sometimes arises is whether the match occurred because of a cross-contamination event in the lab. The concern is particularly acute if the forensic sample and the reference sample were analyzed in the same lab at about the same time. Even if, upon review of the records, lab technicians can find no evidence of a mistake and are therefore completely certain that no mistake was made, the possibility that an undetected contamination event occurred cannot be logically dismissed. Intuitively, the difference between a zero chance of contamination and a very small chance of contamination (e.g., 1-in-1500) seems like the difference between certain evidence of guilt and *nearly* certain evidence of guilt. However, when DNA samples are contemporaneously analyzed in the same lab, the difference between a zero chance of contamination and a very small chance of contamination can instead be the difference between certain evidence of guilt and strong evidence of *innocence*. This article demonstrates that counterintuitive fact by applying a Bayesian statistical analysis to an unusual case where DNA contamination has long been suspected of having occurred despite testimony from lab technicians claiming that the probability of a contamination error was literally zero.

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When an unknown DNA profile found on crime-scene evidence is found to match someone's known DNA profile, a common concern is that it might be a false match.<sup>1</sup> A false match (as that term is used here) occurs when the DNA found on the crime scene evidence belongs to Person A, but the profile is found to match Person B.<sup>2</sup> However, assuming error-free analysis, if the DNA profile from the evidentiary sample is largely intact, the *random match probability* is such that the odds of a false match are extremely remote (e.g., 1 in 10 billion).<sup>3</sup>

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1. William C. Thompson et al., *How the Probability of a False Positive Affects the Value of DNA Evidence*, 48 *J. FORENSIC SCI.* 1, 1 (2003).

2. *Id.*

3. Bruce S. Weir, *The Rarity of DNA Profiles*, 1 *ANNALS APPLIED STAT.* 358, 369 (2007); see WILLIAM C. THOMPSON, *THE POTENTIAL FOR ERROR IN FORENSIC DNA TESTING (AND HOW THAT COMPLICATES THE USE OF DNA DATABASES FOR CRIMINAL IDENTIFICATION)* 6 (2008).

Another concern that arises in some cases is not with that vanishingly unlikely possibility but instead with the possibility that an error occurred while the DNA was being analyzed, such as the mislabeling of evidence samples.<sup>4</sup> For example, if a DNA profile obtained from a drinking glass is mislabeled as having come from a rape kit, the individual who drank from that glass will match and be suspected of having committed the rape. Using a Bayesian analysis, Professors William Thompson, Franco Taroni, and Colin Aitken showed that even a relatively low rate of human error in the lab (e.g., 1-in-10,000 tests) can introduce far more uncertainty about the value of DNA evidence than is implied by an extremely low random match probability.<sup>5</sup> When there is concern that an error like that may have occurred in the lab, a National Research Council (NRC) report argued that the best approach is to retest the relevant samples.<sup>6</sup>

Other potential lab errors, such as cross-contamination, cannot be corrected so easily.<sup>7</sup> As in the case of mislabeling, the problem would not be that the DNA does not belong to the matching individual (i.e., it would not be a false positive). Instead, it would definitely belong to the matching individual. However, appearances notwithstanding, that fact would not indicate that the matching individual was present at the crime scene. Moreover, as noted by Thompson and his coauthors, retesting will not solve the problem.<sup>8</sup> If, for example, the evidentiary sample was contaminated in the lab by the suspect's reference sample, then retesting will only yield the same misleading result no matter how many times it is repeated.

In a case where there is concern about the possibility of laboratory contamination, the competing hypotheses differ from the usual hypotheses, which focus on whether the evidentiary DNA belongs to the matching individual (Hypothesis 1) or to someone else (Hypothesis 2).<sup>9</sup> In the contamination scenario, the competing hypotheses are "the matching individual's DNA was deposited on the evidence at the crime scene at the time the crime was committed" (Hypothesis 1) versus "the matching individual's DNA was deposited on the evidence due to a contamination event in the lab" (Hypothesis 2). This article provides a Bayesian analysis of the contamination scenario that compares these two hypotheses and uses the case of Gary Lieterman—an actual case where lab contamination has long been suspected<sup>10</sup>—to illustrate why such an analysis is important.

As shown later, the importance of the Bayesian analysis is that it demonstrates that if one assumes no chance of lab contamination (i.e., the contamination rate literally equals 0), then the outcome will support Hypothesis 1.

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4. Thompson et al., *supra* note 1, at 1, 2.

5. *Id.* at 1, 5–8.

6. COMM. ON DNA FORENSIC SCI.: AN UPDATE, NAT'L RESEARCH COUNCIL, THE EVALUATION OF FORENSIC DNA EVIDENCE 25 (1996) [hereinafter NRC EVALUATION OF FORENSIC DNA EVIDENCE].

7. Thompson et al., *supra* note 1, at 2.

8. *Id.*

9. *See id.* at 1.

10. *People v. Lieterman*, No. 265821, 2007 WL 2120514, at \*2 (Ct. App. Mich. July 24, 2007).

However, under certain conditions, if one allows for even a small chance of contamination (e.g., the contamination rate equals 1-in-1500, meaning that 99.93% of analyses are error-free), then, counterintuitively, the outcome strongly supports Hypothesis 2 instead. The conditions under which that reversal occurs are rare, but they were present in Gary Leiterman's case.<sup>11</sup> Indeed, Leiterman's case is one of four possible DNA contamination cases highlighted in a National Institute of Justice report entitled *DNA for the Defense Bar*.<sup>12</sup> Here, this case is used to illustrate that the difference between a contamination rate of literally 0 and a contamination rate as small as 1-in-1500 can be the difference between concluding that the matching individual was at the crime scene (and is therefore likely guilty) versus the matching individual is the victim of a contamination event (and is therefore likely innocent).

## I. THE CONVICTION OF GARY LEITERMAN

In 2005, Gary Leiterman was convicted for the March 1969 murder of Jane Mixer, a University of Michigan graduate student, and sentenced to life in prison without the possibility of parole.<sup>13</sup> His conviction was based on a cold case analysis conducted in early 2002 (33 years after the crime) in the Michigan State Police Laboratory.<sup>14</sup> An unknown DNA profile found on Jane Mixer's pantyhose was entered into a DNA database, where it matched the known profile of Gary Leiterman.<sup>15</sup> His known profile had been entered into the database following his arrest for forging a prescription for pain medication in late 2001 at the age of 59.<sup>16</sup> At that time, he provided a reference DNA sample (saliva), which was then analyzed for inclusion in the Michigan DNA database. The match be-

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11. See, e.g., ERIN E. MURPHY, *INSIDE THE CELL: THE DARK SIDE OF FORENSIC DNA* 56–58 (2015); Simon A. Cole & Michael Lynch, *The Social and Legal Construction of Suspects*, 2 ANN. REV. L. & SOC. SCI. 39, 48 (2006); Christopher Halkides & Kimberly Lott, *Presumptive and Confirmatory Blood Tests*, in FORENSIC SCIENCE REFORM: PROTECTING THE INNOCENT 239, 254 (Wendy J. Koen & C. Michael Bowers eds., 2017); Andrea Roth, *Defying DNA: Rethinking the Role of the Jury in an Age of Scientific Proof of Innocence*, 93 B.U. L. REV. 1643, 1676 (2013) [hereinafter Roth, *Defying DNA*]; Andrea Roth, *Safety in Numbers? Deciding When DNA Alone Is Enough to Convict*, 85 N.Y.U. L. REV. 1130, 1143 n.59 (2010) [hereinafter Roth, *Safety in Numbers?*]; Boaz Sangero & Mordechai Halpert, *Why a Conviction Should Not Be Based on a Single Piece of Evidence: A Proposal for Reform*, 48 JURIMETRICS J. 43, 62 n.85 (2007); William C. Thompson, *Tarnish on the 'Gold Standard': Understanding Recent Problems in Forensic DNA Testing*, CHAMPION, Jan.–Feb. 2006, at 10, 14.

12. NAT'L INST. OF JUSTICE, DEP'T OF JUSTICE, *DNA FOR THE DEFENSE BAR* 137 (2012).

13. See GREGORY A. FOURNIER, *TERROR IN YPSILANTI: JOHN NORMAN COLLINS UNMASKED* 447–49 (2016); David J. Krajicek, *The Elmer Fudd Killer*, N.Y. DAILY NEWS (Mar. 25, 2008, 5:52 PM), <http://www.nydailynews.com/news/crime/elmer-fudd-killer-article-1.266919> [https://perma.cc/U8M9-8MVA].

14. John T. Wixted, *Whether Eyewitness Memory or DNA, Contaminated Forensic Evidence Is Unreliable*, ASS'N FOR PSYCHOL. SCI. (Oct. 31, 2016), <https://www.psychologicalscience.org/observer/whether-eyewitness-memory-or-dna-contaminated-forensic-evidence-is-unreliable> [https://perma.cc/ZW78-GM9P].

15. See *id.*; see also MURPHY, *supra* note 11, at 56 (noting that Leiterman's sweat was found on the victim's pantyhose).

16. See MURPHY, *supra* note 11, at 56.

tween Leiterman's known DNA profile and the unknown evidentiary DNA profile from Mixer's pantyhose was the main evidence leading to his conviction in 2005.<sup>17</sup> Other evidence against him included the fact that he owned a .22 caliber gun in 1969 (Mixer was shot in the head with a .22), as well as the fact that a handwriting expert called by the prosecution testified that two words written on a phonebook near where Mixer was abducted were written by him (but a different handwriting expert called by the defense concluded that those words were not written by him).<sup>18</sup> By all accounts, the DNA evidence was decisive. The occurrence of that DNA match changed Leiterman's status from being just one of millions of people who might have committed the crime to being someone who is now serving a life sentence as Jane Mixer's murderer.

As noted above, serious reservations about the DNA evidence in this case have long been expressed.<sup>19</sup> The main reason for these reservations is that there were actually two unknown DNA profiles obtained from the Mixer crime-scene evidence analyzed in 2002.<sup>20</sup> The second unknown profile was from a blood spot taken from Jane Mixer's left hand back in 1969.<sup>21</sup> When that unknown profile was entered into the DNA database, it matched the known profile of John Ruelas, which was added to the database following Ruelas' arrest for the murder of his mother in early 2002.<sup>22</sup> The fact that he murdered his mother would ordinarily make him a strong suspect for the murder of Jane Mixer as well—except that in 1969, Ruelas was only 4 years old.<sup>23</sup> His young age at the time of Mixer's murder obviously ruled him out as a suspect. Still, the notion that a 4-year-old preschooler was not only present at the murder scene but also bleeding on the victim struck many as being entirely implausible and raised the specter of DNA contamination in the lab—especially considering that no connection between Leiterman and Ruelas was ever established despite an exhaustive inquiry.<sup>24</sup>

Reinforcing that concern was the troubling fact that DNA samples from all three individuals—Mixer, Leiterman, and Ruelas—were independently analyzed at about the same time (in early-to-mid 2002) in the Michigan State Police Laboratory.<sup>25</sup> According to laboratory records, Mixer's cold-case evidence was analyzed in March and April 2002; Leiterman's buccal swab arrived at the lab

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17. Krajicek, *supra* note 13.

18. DONALD E. SHELTON, FORENSIC SCIENCE IN COURT: CHALLENGES IN THE TWENTY-FIRST CENTURY 34–35 (2011); see Krajicek, *supra* note 13. An NRC committee found that there “has been only limited research to quantify the reliability and replicability of the practices used by trained document examiners” and that “[t]he scientific basis for handwriting comparisons needs to be strengthened.” COMM. ON IDENTIFYING THE NEEDS OF THE FORENSIC SCI. CMTY., NAT’L RESEARCH COUNCIL OF THE NAT’L ACADS., STRENGTHENING FORENSIC SCIENCE IN THE UNITED STATES: A PATH FORWARD 166–67 (2009). However, the committee also acknowledged there may be some value to handwriting analysis. *Id.* at 167.

19. See sources cited *supra* notes 11–12.

20. *People v. Leiterman*, No. 265821, 2007 WL 2120514, at \*1 (Ct. App. Mich. July 24, 2007); MURPHY, *supra* note 11, at 56.

21. *Leiterman*, 2007 WL 2120514, at \*1.

22. *Id.* at \*3.

23. *Id.*

24. See sources cited *supra* notes 11–12.

25. MURPHY, *supra* note 11, at 57.

in February 2002 and was analyzed in July 2002; crime-scene evidence from the Ruelas murder arrived at the lab in February 2002 and was analyzed in late March of that year.<sup>26</sup> According to the prosecution's theory, Mixer, Leiterman, and Ruelas were together on the night that Mixer was murdered, and between midnight and 3:00 AM, 26-year-old Leiterman was the murderer and 4-year-old Ruelas was a bleeding bystander. Following this narrative, the fact that the DNA samples from all three were together again in early 2002 in the Michigan State Police Laboratory was merely a coincidence.<sup>27</sup> This theory further holds that no contamination occurred, a claim that was emphatically endorsed by lab technicians who testified that there was no possibility of a cross-contamination event in this case.<sup>28</sup> By contrast, the defense's theory proposed that DNA from Leiterman and Ruelas was inadvertently deposited on the Mixer cold case evidence in 2002 due to an undetected contamination event.<sup>29</sup>

The jury agreed with the prosecution's theory in 2005, and an appeals court ruled in 2007 that the trial was properly conducted.<sup>30</sup> Leiterman is currently serving a life sentence without the possibility of parole.<sup>31</sup>

## II. BAYESIAN ANALYSIS

Bayesian inference is a widely accepted "statistical method of inductive reasoning based on the reassessment of competing hypotheses in the presence of new evidence."<sup>32</sup> The statistical analysis presented in this article compares two competing hypotheses: the prosecution's hypothesis that Leiterman's DNA was deposited on Mixer's clothing at the crime scene in 1969 (and is therefore guilty) versus the defense's hypothesis that Leiterman's DNA was deposited on Mixer's clothing in the DNA lab in 2002 (and is therefore innocent). Each hypothesis involves several inputs that bear on the likelihood that the hypothesis is true, and these inputs determine the outcome of the analysis by pointing in favor of guilt or innocence. All of the inputs in our analysis are explicit and have an empirical basis. Most are intuitively sensible, but this does not mean that they are necessarily the best estimates. However, having them explicit makes them easy to challenge (and change) if any seem unreasonable or unduly biased in favor of the prosecution or the defense.

It is important to emphasize at the outset that the main analysis is based on the DNA evidence, not on any other evidence that is potentially relevant to this case. For example, this new statistical analysis does not take into consideration the fact that Leiterman owned a .22 caliber gun in 1969 (which is unrelated to the DNA analysis and points in the direction of Leiterman's guilt) or the fact

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26. *Id.*

27. FOURNIER, *supra* note 13, at 26; *see* Cole & Lynch, *supra* note 11; Thompson, *supra* note 11.

28. Thompson, *supra* note 11.

29. *People v. Leiterman*, No. 265821, 2007 WL 2120514, at \*2 (Ct. App. Mich. July 24, 2007).

30. *Id.* at \*1.

31. *Id.*; Krajicek, *supra* note 13.

32. *Bayesian Inference*, NATURE, <https://www.nature.com/subjects/bayesian-inference> [<http://perma.cc/MN5B-M7YG>].

that serial killer John Norman Collins was operating in the area at the time of Mixer's murder and killed another University of Michigan graduate student 10 weeks after the Mixer murder, by shooting her in the head with a .22 caliber gun (which is also unrelated to the DNA analysis and points in the direction of Leiterman's innocence).<sup>33</sup> Instead, the main analysis presented here ignores all other non-DNA evidence and makes the assumption that a determination of guilt or innocence hinges on the strength of the DNA evidence.<sup>34</sup> This analysis focuses primarily on the DNA evidence to underscore the counterintuitive point that the difference between a contamination rate of 0 and a contamination rate as small as 1-in-1500 can be the difference between concluding that the matching individual is guilty or innocent.

A key input to our statistical analysis is the probability that a DNA analysis would result in a cross-contamination event. Such events are undoubtedly rare, but they do occur. According to an NRC report on DNA technology, "Laboratory errors happen, even in the best laboratories and even when the analyst is certain that every precaution against error was taken."<sup>35</sup> Similarly, Peter Gill and Amanda Kirkham argue that "it should be recognized that laboratory contamination is impossible to avoid completely but its extent is generally unknown unless proactively assessed—the probability of contamination must always be greater than zero."<sup>36</sup> Yet, at the 2005 Leiterman trial, lab personnel testified that cross contamination was impossible because every precaution had been taken to guard against contamination errors.<sup>37</sup> In terms of this analysis, if the lab technicians were right, it would mean that the estimated contamination rate for the Leiterman case should be set to 0. If that were true, then the only reasonable conclusion would be that the Leiterman's DNA was deposited on Mixer's clothing at the crime scene (consistent with the prosecution's theory).

The statistical analysis presented below returns a probability of 1 that the prosecution's theory is correct when the contamination rate is assumed to be 0 (as it must). However, this analysis also shows that changing the contamination rate from 0 (i.e., 100% error-free) to a value as low as 1-in-1500 (i.e., 99.93% error-free) results in an outcome that completely reverses the conclusion. That is, if the contamination rate is only 1-in-1500, such that 99.93% of analyses are error-free, then the estimated probability is close to 1 that the *defense's* theory is true. Intuitively, by contrast, the difference between 100% of DNA analyses

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33. See FOURNIER, *supra* note 13, at 79–82; Walker Lundy & Tom DeLisle, *A Pattern of Death: Anatomy of 7 Brutal Murders*, DETROIT FREE PRESS, July 28, 1969, at 12-D.

34. The only other slightly incriminating evidence presented at the trial consisted of a decades-old memory of a former roommate according to which Leiterman collected newspapers containing articles about a serial killer suspected of having committed several other murders in the Ann Arbor area during the late 1960s. *People v. Leiterman*, No. 265821, 2007 WL 2120514, at \*2 (Ct. App. Mich. July 24, 2007).

35. COMM. ON DNA TECH. IN FORENSIC SCI., NAT'L RESEARCH COUNCIL, DNA TECHNOLOGY IN FORENSIC SCIENCE 89 (1992).

36. Peter Gill & Amanda Kirkham, *Development of a Simulation Model to Assess the Impact of Contamination in Casework Using STRs*, 49 J. FORENSIC SCI. 485, 491 (2004).

37. See Transcript of Julie French at 169–207, *People v. Leiterman*, No. 04-2017-FC, 2004 WL 5546611 (Mich. Cir. Ct. Jan. 1, 2004) (testimony of Sarah Thivault of the Michigan State Police Forensic Science Division about the procedures to prevent cross-contamination).

being free of cross-contamination errors and 99.93% of DNA analyses being free of cross-contamination errors seems like the difference between certain guilt and *almost* certain guilt. Contrary to that powerful intuition, under some circumstances—such as the circumstances that apply to the Gary Leiterman case—it can be the difference between certain guilt and almost certain innocence (where, in this case, “guilt” refers to having deposited DNA at the crime scene, and “innocence” refers to the DNA having been deposited by accident in the crime lab). The formal analysis is presented next.

### III. THE COMPETING HYPOTHESES AND RELEVANT DATA

Three mutually exclusive and exhaustive hypotheses need to be considered:

- $H_1$ : Leiterman’s DNA was deposited on Mixer’s clothing only at the crime scene in 1969.
- $H_2$ : Leiterman’s DNA was deposited on Mixer’s clothing only in the lab in 2002.
- $H_3$ : Leiterman’s DNA was deposited on Mixer’s clothing at the crime scene in 1969 and in the lab in 2002.

The probability of  $H_3$  is obviously extremely low because it assumes that the person who likely murdered Jane Mixer in 1969 just happened to be the same person whose DNA contaminated her evidence in the lab thirty-three years later, but the hypothesis is included for completeness. The bulk of our analysis focuses on the odds of  $H_1$  versus  $H_2$ , the two plausible hypotheses. However,  $H_3$  is formally considered when we ultimately compute the estimated probability of guilt even though its likelihood is much too low to appreciably affect the end result.

The next step is delineating the relevant data for analysis. According to our reading of the trial transcripts, there were four main observations called “events” ( $E$ ) that comprise the relevant data. The following four events are not disputed by either the defense or the prosecution:

- $E_m$ : There was a definite match (hence, the subscript  $m$ ) between Leiterman’s known profile and the DNA found on Mixer’s pantyhose.<sup>38</sup>
- $E_s$ : The matching DNA on Mixer’s pantyhose was consistent with saliva ( $s$ ). That is, trial testimony from lab technicians indicated that Leiterman’s DNA on Mixer’s pantyhose was not consistent with blood or semen but was consistent with other biological materials, including saliva.<sup>39</sup>

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38. *People v. Leiterman*, No. 265821, 2007 WL 2120514, at \*6 (Ct. App. Mich. July 24, 2007); MURPHY, *supra* note 11, at 56.

39. *Leiterman*, 2007 WL 2120514, at \*6.

- $E_e$ : Leiterman's DNA was exclusively ( $e$ ) found on Mixer's pantyhose (i.e., there was no DNA from Mixer herself).<sup>40</sup>
- $E_c$ : Leiterman's reference sample following his arrest in 2001 (a saliva sample) and Mixer's cold-case crime evidence from 1969 were in the lab contemporaneously ( $c$ ) in the first 6 months of 2002, though they were not analyzed on the same days.<sup>41</sup>

A key component of this analysis consists of estimating the probability of each event assuming  $H_1$  and, separately, assuming  $H_2$ . We make the assumption that the four events are statistically independent. Therefore, whether  $H = H_1$  or  $H = H_2$ :

$$P(D|H) = P(E_m|H) \times P(E_s|H) \times P(E_e|H) \times P(E_c|H),$$

where  $P$  indicates the probability of the event in parentheses, and the vertical bar indicates a conditional probability. Thus, for example,  $P(D|H_1)$  represents the probability of the data given Hypothesis 1.

By the Law of Conditional Probability:<sup>42</sup>

$$\frac{P(H_1|D)}{P(H_2|D)} = \frac{P(D|H_1)}{P(D|H_2)} \times \frac{P(H_1)}{P(H_2)}$$

This equation is Bayes' rule in odds form. The ratio on the left-hand side, the relative probability that the hypotheses are true in light of the relevant data, is the target of interest. It is called the "posterior odds."<sup>43</sup> The term on the far right is the prior odds, the probability that the hypotheses are true before any data are observed.<sup>44</sup> The middle term is the likelihood ratio, and it describes how the data have influenced the probabilities that the hypotheses are true.<sup>45</sup> It is helpful here to restate Bayes rule in verbal form:

$$\text{posterior odds} = \text{likelihood ratio} \times \text{prior odds}$$

The posterior odds of  $H_1$  versus  $H_2$  is the target of interest for assessing the guilt of Leiterman. For example, if the outcome of the analysis yields a posterior odds of 100 / 1, it would mean that the odds are 100-to-1 in favor of  $H_1$  (the prosecution's hypothesis) given the data. Both the prior odds and the likelihood ratio must be calculated before the posterior odds of guilt can be computed, which can then be translated into a posterior probability of guilt.

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40. *Leiterman*, 2007 WL 2120514, at \*6

41. MURPHY, *supra* note 11, at 57.

42. See CHARLES M. GRINSTEAD & J. LAURI SNELL, GRINSTEAD AND SNELL'S INTRODUCTION TO PROBABILITY 134–35 (2006).

43. Joe Felsenstein, Adjunct Professor of Comput. Sci. & Statistics, Univ. of Wash., Likelihood and Bayesian Inference (2011), at 4, <http://evolution.gs.washington.edu/gs560/2011/lecture7.pdf>.

44. *Id.*

45. *Id.*



#### IV. THE PRIOR ODDS

The prior odds are equal to  $P(H_1) / P(H_2)$ .  $P(H_1)$  is the probability that—before anything is known about the identity of the person whose DNA was deposited on Mixer’s clothing in 1969—it would turn out to be Leiterman. The prior probability that the matching individual would turn out to be Leiterman is set to  $P(H_1) = 1 / N$ , where  $N$  is the number of people who would have been suspected of having deposited their DNA at the crime scene if their DNA matched the unknown DNA profile on the Mixer crime scene evidence and nothing else was known about their personal history (e.g., about potential alibis).

To estimate  $N$ , the first step is to set the radius around Ann Arbor, Michigan, (where Mixer lived) within which plausible candidates for having been at the crime scene resided.<sup>46</sup> We set this radius based on the prosecution’s interpretation of the DNA evidence. Specifically, both Leiterman and Ruelas were judged by the prosecution to have deposited their DNA at the crime scene.<sup>47</sup> Ruelas lived in Detroit in 1969 about 40 miles from Ann Arbor, and Leiterman lived in the outskirts of Detroit about 20 miles from Ann Arbor.<sup>48</sup> We therefore set a radius of 40 miles around Ann Arbor, which just barely includes where both Leiterman and Ruelas lived in 1969. At a minimum, anyone living in that region in 1969 whose DNA was found on Mixer’s evidence would have been judged by the prosecution to have been at the crime scene. It seems reasonable to suppose that the radius (and the corresponding population) would actually be considerably larger than that, but we use a 40-mile radius as a conservative estimate. It is a conservative estimate (favoring the prosecution) because if the DNA had matched an adult male who lived as far away as 80 miles from the crime scene in 1969, it seems likely he would have been regarded as a viable suspect.

Today, the population living within a 40-mile radius of Ann Arbor is approximately 4.4 million people.<sup>49</sup> According to the U.S. Census Bureau, the entire population of Michigan in 1969 was ~8.8 million people,<sup>50</sup> which is about 89 percent of what it is today.<sup>51</sup> Thus, a reasonable value for this parameter would be  $N = .89 \times 4.4$  million people  $\approx 4$  million people. Keep in mind that

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46. See Kajicek, *supra* note 13.

47. See MURPHY, *supra* note 11, at 56.

48. *Deadly Ride: After 30 Years, a Suspect Is Charged in Coed’s Murder*, CBS NEWS: 48 HOURS (Nov. 22, 2005), <https://www.cbsnews.com/news/deadly-ride/5/> [<https://perma.cc/U94F-MNLW>].

49. *Find Population on Map*, FREE MAP TOOLS, <https://www.freemaptools.com/find-population.htm> (last visited Jan. 4, 2018) (estimating population based on a defined area of a forty-mile radius of Ann Arbor, Michigan).

50. BUREAU OF THE CENSUS, U.S. DEP’T OF COMMERCE, SER. P-25 NO. 437 POPULATION ESTIMATES AND PROJECTIONS 11 (1970) (providing that Michigan’s provisional population estimate for 1969 was 8,766,000).

51. POPULATION DIV., U.S. CENSUS BUREAU, TABLE 1. ANNUAL ESTIMATES OF THE RESIDENT POPULATION FOR THE UNITED STATES, REGIONS, STATES, AND PUERTO RICO: APRIL 1, 2010 TO JULY 1, 2017, at 1 (2017), <https://www.census.gov/data/tables/2017/demo/popest/nation-total.html> [<https://perma.cc/Z828-3Y56>] (providing that Michigan’s current population estimate is 9,962,311). Download the table from a hyperlink provided on the web page.

this is an estimate of the number of people who, before the fact, would have been judged by the prosecution to have deposited their DNA at the crime scene had their DNA been found on Mixer's crime scene evidence in 2002 (as was true of both Leiterman and Ruelas).

The ultimate question is whether Leiterman murdered Mixer. Thus, if we assume that  $H_1$  is that Leiterman deposited his DNA at the crime scene *because he murdered her*, then many of these four million people would be excluded for the same reason that Ruelas was excluded as being the murderer (even though he was not excluded as having deposited his DNA at the crime scene). To be considered as a plausible suspect for the murder, it seems likely that the individual whose DNA was found on the Mixer evidence would have had to have been an adult male living in the region at the time of the murder. Approximately 50 percent of the Michigan population is male, and approximately 50 percent of Michigan males are between the ages of 18 and 59.<sup>52</sup> Thus, we can reasonably set  $N = .50 \times .50 \times 4 \text{ million} = 1 \text{ million}$ . This is an estimate of the number of people who, before anything more than age and sex was known about them, would have been judged by the prosecution to have deposited their DNA at the crime scene had their DNA been found on Mixer's crime scene evidence in 2002 and who would have been suspected of having murdered her (as Leiterman was, but Ruelas was not). That is, in the absence of any further individuating evidence, each of those adult men had about a one in a million chance of being the killer.

A disadvantage to this approach is that  $N$  could be made arbitrarily large or small to suit one's ends. For example, to minimize  $N$ , favoring the prosecution, one could more narrowly define it to be the number of males between the ages of 17–44 in 1969 who lived within 20 miles of the crime (as Leiterman did) and who were not married (as Leiterman was not).<sup>53</sup> However, in Appendix A, an alternative approach to estimating  $N$  is described, which relies solely on DNA database search statistics (not on any defense- or prosecution-oriented assumptions), and arrives at approximately the same value as above (~1 million).<sup>54</sup> In addition, although  $N = 1 \text{ million}$  is used, as shown later, similar conclusions are reached even when much lower values of  $N$  are assumed (i.e., even when the value is implausibly biased to favor the prosecution's theory). Consequently, for the main analysis,  $P(H_1) = 1 / 1 \text{ million} = 10^{-6}$ .

Next, the prior probability of the defense's hypothesis,  $P(H_2)$ , can be estimated. This is the probability that, before anything is known about who matched, Leiterman would be the one whose DNA was deposited on Mixer's clothing. What is the probability that, given a contamination event, the person involved would be Gary Leiterman? According to Kyle Jen, Senior Fiscal Ana-

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52. *Current Population Demographics and Statistics for Michigan by Age, Gender, and Race*, SUBURBAN STATS, <https://suburbanstats.org/population/how-many-people-live-in-michigan> [<https://perma.cc/FP6D-AZN5>].

53. See Krajicek, *supra* note 13.

54. See *infra* Appendix A.

lyst with the Michigan House Fiscal Agency, 9933 reference samples were analyzed in Michigan lab in 2002.<sup>55</sup> We therefore assume that approximately half that number ( $n \approx 5000$ ) were analyzed contemporaneously with the Mixer evidence in the first half of 2002. Thus, the probability that Leiterman is the one of these 5000 is  $P(H_2) = 1 / n = 1 / 5000 = 2 \times 10^{-4}$ . Note that this number could be reasonably set to a smaller value because it has not been reduced by limiting it to males of the appropriate age. However, reducing it would favor  $H_2$  (the defense's theory), so we conservatively assume that all 5000 would be plausible suspects even though the age (too young) and sex (female) of at least some would exclude them.

Having estimated  $P(H_1)$  and  $P(H_2)$ , we are in a position to estimate the prior odds,  $P(H_1) / P(H_2)$ . With  $P(H_1) = 10^{-6}$  and  $P(H_2) = 2 \times 10^{-4}$ , the prior odds come to  $P(H_1) / P(H_2) = 0.005$  (i.e., 1-in-200). This value means that, before knowing anything else about the evidence considered below, a DNA match would imply that if Leiterman matched, it was 200 times more likely that his DNA was deposited in the lab due to a contamination event than at the crime scene while murdering Jane Mixer. Hence, the prosecution has a large burden. But if the probability of contamination is 0, the DNA match overcomes this burden. The question of interest is whether the same is true when the probability of contamination is close to 0 (e.g., 1-in-1500) but not exactly 0.

## V. THE LIKELIHOOD RATIO

Next, the parameters needed to compute the likelihood ratio must be estimated. To do so, the probability of observing each of the four relevant events given  $H_1$  and, separately, given  $H_2$ , is computed.

### A. $E_m$ (the DNA Match)

Under what conditions would a match occur assuming  $H_1$ ? This is the probability that Mixer's murderer would deposit his DNA on her clothing in the first place ( $p_1$ ) times the probability that the DNA would survive 33 years of storage ( $p_2$ ). That is,  $P(E_m|H_1) = p_1 \times p_2$ . We first estimate  $p_1$ .

According to Jen, in the Michigan State Police Laboratory circa 2002, only half of the recent (i.e., not cold case) forensic evidence received by the lab yielded a DNA profile that was sufficiently intact to enter into a database.<sup>56</sup> This means that in only fifty-percent of criminal cases analyzed someone—presumably the perpetrator in most cases—left an identifiable DNA profile on the evidence. A similar result was observed in a prospective study of DNA collected from burglary crime scenes.<sup>57</sup> Out of 1079 cases, the researchers found that 54.7

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55. See KYLE I. JEN, MICH. HOUSE FISCAL AGENCY, MICHIGAN'S FORENSIC DNA DATABASE 5 (2003), [www.house.mi.gov/hfa/Archives/PDF/dna.pdf](http://www.house.mi.gov/hfa/Archives/PDF/dna.pdf) [<https://perma.cc/NPL8-E2UK>].

56. *Id.*

57. See John K. Roman et al., *The DNA Field Experiment: A Randomized Trial of the Cost-Effectiveness of Using DNA to Solve Property Crimes*, 5 J. EXPERIMENTAL CRIMINOLOGY 345, 356 (2009).

percent resulted in the generation of a profile that was sufficiently intact to search a database.<sup>58</sup> Thus, we assume that  $p_1 = .50$ .

Consider next the probability that DNA on crime scene evidence would survive 33 years of storage in the first place ( $p_2$ ). A news story published in 2011 about a cold case unit in Iowa provides some guidance on this issue.<sup>59</sup> The story indicated that “the unit collected data on 150 unsolved cases going back to the 1960s. The investigators prioritized 50 of them and conducted DNA analysis on 2,018 pieces of evidence. They developed profiles in 23 cases and checked 11 of those profiles against the nationwide DNA database.”<sup>60</sup> In other words, 11 of the 50 cold cases subjected to DNA analysis yielded a DNA profile sufficiently intact to enter into a database ( $11 / 50 = .22$ ), as happened in the Mixer cold case investigation. For reasons described above, we would expect only half the cases to yield a usable DNA profile even if it were a recent case involving little to no degradation (because DNA was found on recent evidence only 50 percent of the time). It therefore follows that for the ~50 percent of cases in which DNA was initially deposited on the evidence it degraded in storage about half the time (i.e.,  $.50 \times .50 = .25$ ), so the expected value is close to the observed value of .22. Thus, it can be assumed that DNA would survive cold-case, long-term storage with a probability of .50 ( $p_2 = .50$ ), which seems like an intuitively sensible value. If anything, it is probably conservative (favoring an outcome of guilt rather than innocence) because the cases investigated by the Iowa cold case unit were probably less than 33 years old on average. For the analysis presented here, the probability of DNA being deposited by the perpetrator while committing the crime and surviving 33 years of storage assuming  $H_1$  is set to  $P(E_m|H_1) = p_1 \times p_2 = .25$ .

The probability of a match given  $H_2$ ,  $P(E_m|H_2)$  is the probability that Leitterman’s DNA was deposited on Mixer’s clothing (and would therefore match) due to a contamination event in the lab. The probability of a contamination event is initially set to  $1 / 1500$  in the main analysis based on recent research by Professor Ate Kloosterman and his colleagues.<sup>61</sup> They reviewed lab records for 472,127 DNA analyses conducted at the Human Biological Traces Department of the Netherlands Forensic Institute (NFI) over the years 2008–2012.<sup>62</sup> Kloosterman and his colleagues found that cross-contamination events with other samples in the lab occurred 311 times in the 2008–2012 time period, which is to say that such an event occurred in  $311 / 472,127 \approx 1 / 1500$  analyses.<sup>63</sup> Note that these are *detected* cross-contamination events, most of which were caught

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58. *Id.* at 357.

59. Mike Wiser, *Time, Money Running Out for State’s Cold Cases*, GLOBE GAZETTE (June 25, 2011), [http://globegazette.com/news/time-money-running-out-for-state-s-cold-cases/article\\_d99154ce-9fa9-11e0-a3e3-001cc4c002e0.html](http://globegazette.com/news/time-money-running-out-for-state-s-cold-cases/article_d99154ce-9fa9-11e0-a3e3-001cc4c002e0.html) [https://perma.cc/7X7N-QMLY].

60. *Id.*

61. See Ate Kloosterman et al., *Error Rates in Forensic DNA Analysis: Definition, Numbers, Impact and Communication*, 12 FORENSIC SCI. INT’L: GENETICS 77, 80–81 (2014).

62. *Id.* at 78, 79 tbl. 2.

63. *Id.* at 80 tbl. 5.

by the NFI quality control systems.<sup>64</sup> Our actual interest is in estimating the rate of *undetected* cross-contamination events. According to  $H_2$ , this is the type of event that occurred in the Leiterman case.

There is no easy way to precisely determine how often undetected events occur because it would require a technique like blind testing of lab technicians, which, given the rarity of contamination events, would be a prohibitively expensive undertaking.<sup>65</sup> Nevertheless, the results reported by Kloosterman and his coauthors indicate that cross-contamination events do occur at a low rate.<sup>66</sup> This finding underscores a point stressed in the 1996 NRC report: “No amount of effort and improved technology can reduce the error rate to zero.”<sup>67</sup> It seems reasonable to suppose that the error rate reported by Kloosterman and his colleagues gives us a ballpark estimate of the value of *undetected* cross-contamination events.<sup>68</sup> Based on the assumption that the rates of detected and undetected contamination are similar, we initially set  $P(E_m|H_2)$  to  $1 / 1500 = 6.67 \times 10^{-4}$ . Later, we consider how the results change when its value is set to within an order of magnitude of that starting value (i.e.,  $1 / 150$  and  $1 / 15,000$ ).

## **B. $E_s$ (the DNA Is Consistent with Saliva)**

The second relevant event ( $E_s$ ) is that the DNA on Mixer’s pantyhose was not consistent with blood or semen but was consistent with saliva. Consider first  $P(E_s|H_1)$ , which is the probability that  $E_s$  would be observed given  $H_1$  (the prosecution’s theory). If Leiterman killed Mixer ( $H_1$ ), the DNA on her pantyhose did not have to be consistent with saliva. Had the DNA been from semen, for example, a match to his known reference sample (from saliva) would still have occurred. Thus, according to  $H_1$ , consistency with saliva was a coincidence that occurred by chance with some probability. We estimate the probability of such a coincidence using data reported by Professor Theodore Cross and his coauthors.<sup>69</sup>

Cross and his colleagues examined forensic evidence from a large number of sexual assault cases in Massachusetts.<sup>70</sup> The Mixer murder did not involve rape but did appear to have a sexual motivation given that her pantyhose had been pulled down.<sup>71</sup> Their Table 5.15 shows the percentage of cases in which

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64. *See id.* at 78.

65. Thompson et al., *supra* note 1, at 7.

66. *See* Kloosterman et al., *supra* note 61, at 80–81 tbls. 5–7.

67. NRC EVALUATION OF FORENSIC DNA EVIDENCE, *supra* note 6, at 4.

68. *See* Kloosterman et al., *supra* note 61.

69. *See generally* THEODORE P. CROSS ET AL., FORENSIC EVIDENCE AND CRIMINAL JUSTICE OUTCOMES IN A STATEWIDE SAMPLE OF SEXUAL ASSAULT CASES (2014), <https://www.ncjrs.gov/pdffiles1/nij/grants/248254.pdf> [<https://perma.cc/YQ8T-EAG7>].

70. *Id.* at E-6 to -7.

71. Transcript of Record at 22, *People v. Leiterman*, No. 04-2017-FC (Mich. Cir. Ct. July 22, 2005) (closing statement of Steven Hiller, Washtenaw County Deputy Chief Assistant Prosecuting Attorney); EARL JAMES, CATCHING SERIAL KILLERS: LEARNING FROM PAST SERIAL MURDER INVESTIGATIONS 34 (1991).

various kinds of biological evidence were detected.<sup>72</sup> According to that table, blood was detected 25.5% of the time, saliva 34.6% of the time, semen 61.4% of the time, and other biological materials 41.2% of the time.<sup>73</sup> These percentages add up to more than 100% because a single case can yield more than one type of biological evidence. From the data given, one can reasonably infer that when only one kind of biological evidence is found, such evidence will not be consistent with blood or semen but will be consistent with saliva and other biological materials (as was true of the Leiterman case) approximately 50% of the time. This is not an exact estimate given that the biological sources may not be independent. Moreover, these data come from sexual assault cases, whereas (as noted earlier) the Mixer case may have been a case of aborted sexual assault. Nevertheless, it is a relevant, empirically-based estimate. Thus,  $P(E_s|H_1)$  was set to .50.

Next consider  $P(E_s|H_1)$ , which is the probability that  $E_s$  (the forensic DNA was consistent with saliva) would be observed given  $H_2$  (the defense's theory).  $H_2$  holds that saliva from Leiterman's buccal swab is what contaminated the Mixer evidence. In other words, according to  $H_2$ , the probability is 1 that the DNA found on Mixer's pantyhose would be consistent with saliva. Thus,  $P(E_s|H_2) = 1$ . This setting ensures that had the DNA come from a source not consistent with saliva (such as blood or semen), the defense's contamination theory would be conclusively ruled out.

### C. $E_e$ (Leiterman's DNA, but Not Mixer's DNA, Found on Mixer's Clothing)

The third relevant event ( $E_e$ ) is the observation that DNA on the pantyhose came exclusively from Leiterman, with no detectable DNA from Mixer herself.<sup>74</sup> Intuition suggests that the opposite would usually be true (i.e., that there would be more of Mixer's DNA on her own pantyhose than Leiterman's DNA).

Assuming that the pantyhose were stored in such a way that the DNA would survive for 33 years (which it was according to  $H_1$ , given that Leiterman's DNA survived), what is the probability that more of Leiterman's DNA was deposited on Mixer's pantyhose than from Mixer herself? Two recent research articles offer some guidance about the probability of finding the outcome observed in the Leiterman/Mixer case, namely, a measurable amount of DNA from Leiterman (the "toucher") on Mixer's pantyhose but with no measurable DNA from Mixer (the "wearer").<sup>75</sup> Professor Michelle Breathnach and her coauthors investigated "the frequency of detection of DNA from wearer, toucher or others when

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72. CROSS ET AL., *supra* note 69, at 91 tbl. 5.15.

73. *Id.* at 91 tbl. 5.15 (showing average of "Non-SANE" and "SANE" columns).

74. See *People v. Leiterman*, No. 265821, 2007 WL 2120514, at \*6 (Ct. App. Mich. July 24, 2007).

75. See Michelle Breathnach et al., *Probability of Detection of DNA Deposited by Habitual Wearer and/or the Second Individual Who Touched the Garment*, 20 FORENSIC SCI. INT'L: GENETICS 53–60 (2016); M. van den Berge et al., *Prevalence of Human Cell Material: DNA and*

individuals wore and handled worn garments under normal circumstances,” and reported that “[t]oucher and no wearer was observed in 15% of reportable samples.”<sup>76</sup> This suggests that a reasonable setting for  $P(E_e|H_1)$  might be .15.

Similarly, Professor Margreet van den Berge and her coauthors investigated an activity scenario in which the toucher grabbed the trouser leg ankles of the wearer and dragged that individual for one minute.<sup>77</sup> In 48 such cases, their Figure 4C indicates that an outcome somewhat similar to that observed in the Leiterman/Mixer case (that is, DNA from the toucher coupled with virtually no DNA from the wearer) occurred in approximately seven cases.<sup>78</sup> More specifically, in 10 of the 48 analyses, the toucher (called the “grabber” in that study) was the major contributor, and of those 10, 7 exhibited very small amounts of DNA from the wearer.<sup>79</sup> Once again, that outcome was observed with a probability of  $7 / 48 = .15$ . Therefore we set  $P(E_e|H_1)$  to .15. Again, this is not an exact value, but it is reasonable and intuitive.

According to  $H_2$  (the defense’s hypothesis), how likely is it that Leiterman’s DNA would be detected on Mixer’s pantyhose, but her own DNA would not be detected?  $H_2$  implies that Leiterman’s DNA was recently deposited from his buccal swab and was detected for that reason. Thus, under this hypothesis,  $P(E_e|H_2)$  reduces the probability that Mixer’s DNA would not be detected at all, as it was not.

Unlike  $H_1$ ,  $H_2$  does not imply that DNA deposited at the crime scene survived 33 years of storage. Thus,  $P(E_e|H_2)$  is equal to the probability that Mixer’s DNA was not deposited on her evidence in the first place ( $p_3$ ) plus the probability that if it were deposited in the first place ( $1 - p_3$ ) it would degrade to the point of being undetectable after 33 years of storage ( $p_4$ ). That is,  $P(E_e|H_2) = p_3 + (1 - p_3) \times p_4$ . From the toucher-wearer studies discussed above, we estimate that  $p_3 \approx .15$  (i.e., wearers sometimes do not leave a trace of DNA on their own clothing). Earlier, we estimated that the perpetrator’s DNA would survive 33 years of storage with a probability of  $p_2 = .50$ . Assuming the same is true for a victim’s DNA on her own clothing,  $p_4 = .50$ . Thus,  $P(E_e|H_2) = .15 + (1 - .15) \times .50 = .575$ .

#### **D. $E_c$ (Contemporaneous DNA Analyses)**

The fourth relevant event ( $E_c$ ) is that Leiterman’s DNA happened to be in the lab during the same 6-month period in 2002 that the Mixer evidence was in the lab and being analyzed. According to  $H_1$  (the prosecution’s theory), the con-

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*RNA Profiling of Public and Private Objects and After Activity Scenarios*, 21 FORENSIC SCI. INT’L: GENETICS 81–89 (2016).

76. Breathnach et al., *supra* note 75, at 53, 59.

77. Van den Berge et al., *supra* note 75, at 83.

78. *Id.* at 87 fig. 4C.

79. *Id.* For the purposes of this examination, the amount of DNA from the wearer was judged to be very small when it was estimated to be 0.25 ng or less.

temporaneous analysis of the Leiterman and Mixer evidence is simply a coincidence that occurred with some probability less than 1 because Leiterman's buccal swab could have been analyzed at any time and a match to the DNA on Mixer's pantyhose would still have occurred. According to Jen's report, approximately 42,000 DNA analyses had been performed by the Michigan State Police Laboratory and entered into the database by the end of 2003.<sup>80</sup> Leiterman's known profile was added some time in 2004,<sup>81</sup> at which point the database likely had even more entries, but we conservatively assume a database size of 42,000 (if anything, favoring an outcome of guilt). Of all of the profiles in the database at the time Leiterman matched in 2004, what is the probability that the matching profile would have been contemporaneously analyzed with the Mixer evidence in the first 6 months of 2002? Under  $H_1$ , this outcome would occur by chance with a probability of  $5000 / 42,000 = .119$ .<sup>82</sup> Thus, we set  $P(E_c|H_1)$  to .119.

Note that this is a conservative value for another reason as well. Just as  $H_1$  does not require that the Leiterman and Mixer analyses occur at the same time (which happened coincidentally with probability .12), it also does not require that the analyses happen in same *place*. For example, imagine that Leiterman murdered Mixer in 1969 (in accordance with  $H_1$ ) and then moved to California shortly thereafter, where, at the age of 61 in 2001, he forged a prescription, was arrested, and had his known DNA profile entered in the California Combined DNA Index System (CODIS) database. When the unknown DNA profile was taken from Mixer's pantyhose in the Michigan State Police Crime lab in 2002 and entered into the Michigan DNA database, it would not have yielded a match. However, the random match probability for Caucasians with this unknown profile was 1-in-170 trillion.<sup>83</sup> That value far exceeds the criterion for entering the profile into the national (federal) CODIS database, which requires a random match probability at least as low as 1-in-10 million.<sup>84</sup> Following a failed search of the Michigan CODIS database, had the unknown profile from the Mixer evidence been entered into the federal database, it would have matched the known profile of Gary Leiterman. He would then have been discovered to have lived in Michigan in 1969, and he would become a very strong suspect. And rightly so. Under these conditions, the possibility of contamination would be negligible (because, under this scenario, his buccal swab would have been analyzed in a California lab in early 2002, whereas the Mixer evidence was analyzed in a Michigan lab in early 2002). The point is that  $H_1$  (the prosecution's theory) does not require that the analyses were performed at either the same time *or* the same

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80. See JEN, *supra* note 55.

81. John Jefferson, *Cold Hits Meet Cold Facts: Are DNA Matches Infallible?*, TRANSCRIPT, Spring 2008, at 30, <https://www.law.berkeley.edu/files/Transcript.Spring08.Weblow.pdf> [<https://perma.cc/L2L5-SBB9>].

82. See JEN, *supra* note 55.

83. Transcript of Record, *supra* note 71, at 120 (witness testimony of Dr. Stephan Milligan).

84. *Frequently Asked Questions on CODIS and NDIS*, FED. BUREAU INVESTIGATION, <https://www.fbi.gov/services/laboratory/biometric-analysis/codis/codis-and-ndis-fact-sheet> [<https://perma.cc/BKZ3-3P48>].



place. The coincidence factor of .119 used in our analysis only takes into account coincidental timing, not the additional fact of coincidental location (which would justify the use of an even lower value). Thus, the value of  $P(E_4|H_1) = .119$  is, if anything, biased in favor of an outcome of guilt.

Under  $H_2$ , contemporaneous analysis in the same lab is a requirement (i.e., it had to happen for contamination to have happened). Thus,  $P(E_c|H_2) = 1$ .

Putting all of these estimates together, the numerator of the likelihood ratio is

$$P(D|H_1) = P(E_m|H_1) \times P(E_s|H_1) \times P(E_e|H_1) \times P(E_c|H_1) = .25 \times .50 \\ \times .15 \times .119 = .0022313.$$

Likewise, the denominator of the likelihood ratio is

$$P(D|H_2) = P(E_m|H_2) \times P(E_s|H_2) \times P(E_e|H_2) \times P(E_c|H_2) = .000667 \times 1 \\ \times .575 \times 1 = .0003835.$$

That is,  $P(D|H_1) = .0022313$  and  $P(D|H_2) = .0003835$ . Thus, the likelihood ratio is equal to  $.0022313 / .0003835 = 5.82$ . This means it is approximately 6 times more likely that a DNA match would occur if  $H_1$  were true compared to if  $H_2$  were true.

## **VI. THE POSTERIOR ODDS AND POSTERIOR PROBABILITY**

The posterior odds can now be computed given the above estimates of the prior odds and likelihood ratio. The point is not that the estimates provided above are indisputable. Rather, we assume that most would agree that they are at least reasonable, which allow us to reasonably estimate how the posterior odds change under different assumptions about the rate of contamination in the Michigan State Police Laboratory in 2002.

First, what are the posterior odds if a contamination rate of only 1-in-1500 (i.e., 99.93% of DNA analyses are error free) is assumed? Using this estimated contamination rate, as in the analysis presented above, the likelihood ratio comes to 5.82. When multiplied by the prior odds of .005, the estimated posterior odds of  $H_1$  versus  $H_2$  come to  $P(H_1|D) / P(H_2|D) = .0291$ . In other words, despite the fact that Leiterman's DNA matched the DNA found on Mixer's clothing, the odds of guilt are very low (i.e., the odds of innocence are very high). Table 1 provides a summary of our Bayesian analysis.

**Table 1. A Summary of the Bayesian Analysis of the Gary Leiterman/Jane Mixer Case**

Hypotheses:

- $H_1$ : Leiterman's DNA was deposited on Mixer's clothing only at the crime scene in 1969.
- $H_2$ : Leiterman's DNA was deposited on Mixer's clothing only in the lab in 2002.
- $H_3$ : Leiterman's DNA was deposited on Mixer's clothing at the crime scene and in the lab.

Events:

- $E_m$ : A match between Leiterman's DNA profile and the forensic DNA profile.
- $E_s$ : The forensic DNA sample was consistent with saliva.
- $E_e$ : Leiterman's DNA was exclusively found on Mixer's clothing.
- $E_c$ : Leiterman's reference DNA sample and Mixer's forensic evidence contemporaneously analyzed.

Prior Odds ( $H_1$  vs.  $H_2$ ):

- $P(H_1) = 1 / 1 \text{ million} = 10^{-6}$
- $P(H_2) = 1 / 5000 = 2 \times 10^{-4}$
- $P(H_1) / P(H_2) = .005$  (i.e.,  $1 / 200$ )

Likelihood Ratio:

- $P(E_m|H_1) = .25$  (Leiterman's DNA left at scene and survives 33 years, per  $H_1$ )
- $P(E_m|H_2) = 1 / 1500 = 6.67 \times 10^{-4}$  (cross-contamination rate in the lab, per  $H_2$ )
- $P(E_s|H_1) = .50$
- $P(E_s|H_2) = 1$
- $P(E_e|H_1) = .15$
- $P(E_e|H_2) = .575$
- $P(E_c|H_1) = 5000 / 42,000 = .119$
- $P(E_c|H_2) = 1$
- $P(D|H_1) = P(E_m|H_1) \times P(E_s|H_1) \times P(E_e|H_1) \times P(E_c|H_1) = .0022313$
- $P(D|H_2) = P(E_m|H_2) \times P(E_s|H_2) \times P(E_e|H_2) \times P(E_c|H_2) = .0003835$
- $P(D|H_1) / P(D|H_2) = 5.82$

Posterior Odds of Guilt:

- $P(H_1|D) / P(H_2|D) = P(H_1) / P(H_2) \times P(D|H_1) / P(D|H_2) = .0291$

Posterior Probability of Guilt (Appendix B):

- $P(H_3|D) \approx 0$
- $P(\text{Guilt}) = P(H_1|D) + P(H_3|D) \approx P(H_1|D)$
- $P(\text{Guilt}) \approx .0283$

To convert the posterior odds of guilt into a posterior *probability* of guilt, we need to consider the mutually exclusive and exhaustive possibilities:  $H_1$ ,  $H_2$ , and  $H_3$ , where  $H_3$  assumes that Leiterman's DNA was deposited both at the crime scene in 1969 *and* in the lab in 2002. In truth, the probability of joint occurrence is so low that it will not affect our calculations in any appreciable way, but it is considered here for the sake of completeness.

The Bayesian analysis that includes a consideration of  $H_3$  is somewhat complicated and is presented in Appendix B. In that analysis, the posterior odds of guilt for  $H_1$  and  $H_2$ —that is,  $P(H_1|D) / P(H_2|D)$ —are exactly the same as those shown in Table 1:  $P(H_1|D) / P(H_2|D) = 0.0291$ . However, we now also show that  $P(H_3|D) / P(H_2|D) = 5.7957 \times 10^{-6}$ . These two equations have three unknowns, namely,  $P(H_1|D)$ ,  $P(H_2|D)$  and  $P(H_3|D)$ . However, because  $H_1$ ,  $H_2$  and  $H_3$  are mutually exclusive and exhaustive, we also know that

$$P(H_1|D) + P(H_2|D) + P(H_3|D) = 1.$$

We now have three equations and three unknowns. Solving for the values of interest yields  $P(H_1|D) = .0283$ ,  $P(H_2|D) = .9717$ , and  $P(H_3|D) = 5.7957 \times 10^{-6}$  (Appendix B). These three values sum to 1 (except for rounding error) because they constitute the set of mutually exclusive and exhaustive possibilities. Based on the DNA evidence, Leiterman would be guilty if either  $H_1$  or  $H_3$  were true. Thus, the posterior probability of guilt,  $P(G|D)$ , is equal to

$$P(G|D) = P(H_1|D) + P(H_3|D) = .0283 + 5.7957 \times 10^{-6} \approx .0283.$$

If the posterior probability of guilt is .0283, it follows that the posterior probability of innocence,  $P(I|D)$ , is  $1 - .0283 = .9717$ .

This analysis yields a completely different result if it is now assumed that the contamination rate is effectively 0 (not 1-in-1500). The lab technicians testified, in fact, that there was literally no chance of contamination in the Leiterman case despite the disconcerting match to Ruelas.<sup>85</sup> If the contamination rate,  $P(E_m|H_2)$ , is set to a value approaching 0, then  $P(D|H_2)$  would approach 0 as well.<sup>86</sup> In that case, the likelihood ratio,  $P(D|H_1) / P(D|H_2)$ , and the posterior odds,  $P(H_1|D) / P(H_2|D)$ , would both approach infinity, which means that the posterior probability of guilt would approach 1.<sup>87</sup> In other words, if a contamination rate approaching 0 is assumed—despite the Ruelas match—the statistical analysis would indicate that Gary Leiterman deposited his DNA on Mixer's clothing in 1969 with probability approaching 1.<sup>88</sup>

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85. See Transcript of Julie French at 169–207, *People v. Leiterman*, No. 04-2017-FC, 2004 WL 5546611 (Mich. Cir. Ct. Jan. 1, 2004) (testimony of Sarah Thivault of the Michigan State Police Forensic Science Division about the procedures to prevent cross-contamination).

86. See Thompson et al., *supra* note 1, at 6.

87. See *id.*

88. See *id.*

The critical point to take away from this analysis is that, contrary to a powerful intuition, it is not the case that 100% error-free DNA analyses implies that  $H_1$  is certain, whereas 99.93% error-free implies that  $H_1$  is *almost* certain. Instead, 100% error-free DNA analyses implies that  $H_1$  is certain (as before), whereas 99.93% error-free implies that  $H_1$  is almost certainly wrong and  $H_2$  is almost certain.

## VII. ALTERNATIVE SCENARIOS

Our central conclusion is reasonably robust to variations in some of the key estimates. For example, setting  $N$  (the number of plausible suspects prior to knowing who matched) to the implausibly low value of 250,000—as if only 250,000 of the 8.8 million Michigan residents in 1969<sup>89</sup> would have become a suspect in the Jane Mixer murder had their DNA matched—only reduces the probability that  $H_2$  is true to .90. Thus, for the analysis to yield any outcome close to favoring  $H_1$ , some other parameter would also have to be changed to an implausible value favoring that outcome.

Setting  $N$  back to a seemingly more reasonable value of 1 million, we can ask about the posterior odds with contamination rates within an order of magnitude of 1-in-1500. Setting the contamination rate to 1-in-150, the posterior probability of  $H_1$  and  $H_2$  come to .003 and .997, respectively. In other words, with a contamination rate that high, a decision based on the DNA evidence alone would indicate almost certain innocence. A contamination rate of 1-in-150 does not seem altogether unreasonable given the widely noted signal that something may have gone wrong in the Michigan State Police crime lab in early 2002 (namely, the disconcerting Ruelas match). Nevertheless, we next assume the opposite. That is, that the lab procedures in early 2002 were so impeccable that only 1-in-15,000 analyses resulted in a cross-contamination event. Setting the contamination rate to 1-in-15,000, the posterior probability of  $H_1$  and  $H_2$  come to .23 and .77, respectively. In other words, even then, the evidence points decidedly in favor of innocence.

What happens if we now bias the analysis in favor of an outcome of guilt to a seemingly absurd degree by assuming both that the number of plausible suspects prior to knowing who matched was only 250,000 (i.e.,  $N = 250,000$ ) and that the undetected cross-contamination rate in early 2002 was an impressive 1-in-15,000 analysis (despite the inexplicable match to Ruelas)? Even then, the probability of guilt is only .54 (i.e., probability of innocence = .46). In other words, to tip the scales ever so slightly in favor of guilt, we need to make multiple absurd assumptions designed to yield an outcome in favor of guilt.

The analysis summarized in Table 1 illustrates the essence of our Bayesian statistical analysis, but it is incomplete in one important respect. Specifically, the inputs to the analysis consisted of fixed estimates. For example,  $P(E_s|H_1)$  (the probability that the DNA found on Mixer's pantyhose would be consistent with saliva) was fixed at .50. Although the estimates used in our analysis were all based on prior knowledge, their exact values are unknown, which

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89. BUREAU OF THE CENSUS, *supra* note 50.

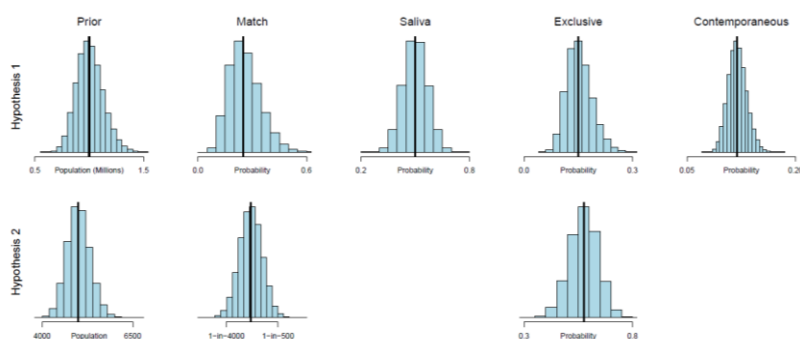
raises the question: How would our conclusions change if we introduced variability into those inputs to reflect uncertainty as to their true values? Accommodating uncertainty in the inputs is a fairly technical matter that need not be considered by readers who are mainly interested in the take-home message. However, the analysis would be incomplete if we did not investigate the degree to which the main conclusion is dependent on the exact values for the inputs, so we turn now to a consideration of that issue.

### VIII. MODELING UNCERTAINTY

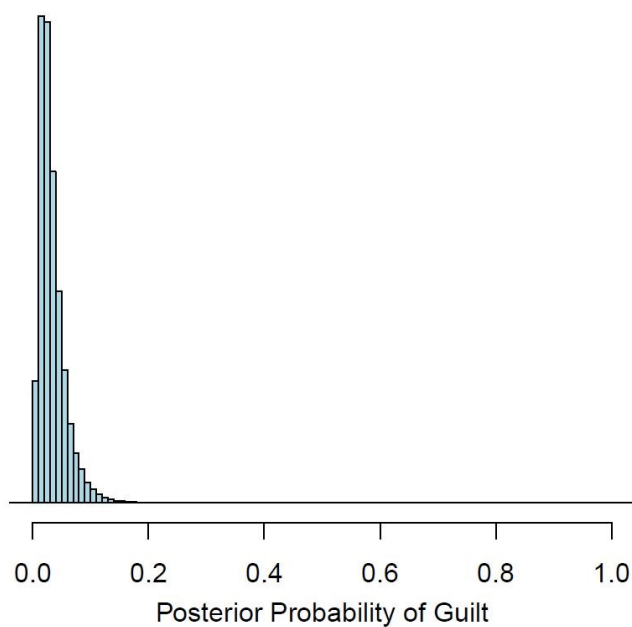
To model uncertainty in the inputs of our Bayesian analysis, we created probability distributions for our estimates of  $P(H_1)$ ,  $P(E_m|H_1)$ ,  $P(E_s|H_1)$ ,  $P(E_e|H_1)$ ,  $P(E_c|H_1)$ ,  $P(H_2)$ ,  $P(E_m|H_2)$ , and  $P(E_e|H_2)$ . The only estimates for which no variability was introduced are the ones that, according to  $H_2$ , must equal 1. More specifically, if  $H_2$  is true, then the DNA on Mixer's clothing had to be consistent with saliva, so  $P(E_s|H_2)$  remained fixed at 1. Similarly, according to  $H_2$ , the samples had to be contemporaneously analyzed in the lab, so  $P(E_c|H_2)$  remained fixed at 1. We introduced variability in all other inputs by treating them as a distribution rather than as a constant.

Figure 1 shows the distributions used for each input in this analysis. The top row shows the distributions for the inputs associated with  $H_1$ , with each panel corresponding to 1 of the 5 inputs:  $P(H_1)$ ,  $P(E_m|H_1)$ ,  $P(E_s|H_1)$ ,  $P(E_e|H_1)$ , and  $P(E_c|H_1)$ . The bottom row shows the corresponding distributions for the inputs associated with  $H_2$ . The distributions for  $P(E_s|H_2)$  and  $P(E_c|H_2)$  are missing in the bottom row because those values were fixed at 1, as required by  $H_2$ . For each distribution, the black vertical line corresponds to the fixed value that was used previously in the main analysis (i.e., the black vertical lines correspond to the values shown in Table 1). For example, the top left panel of Figure 1 shows the distribution of the value of  $N$  that corresponds to the uncertainty in  $P(H_1)$ . For the main analysis,  $N$  was fixed at 1 million, which is the value represented by the black vertical line. Other uncertainties are expressed directly as distributions on probabilities or distributions on rates, as in the cross-contamination rate. The guiding principle in setting these variabilities was to be liberal, that is, to try to err on the side of too much variability. The ranges subtended seem to us to be the maximal plausible ranges.

**Figure 1. Distributions of Inputs Used to Compute the Posterior Distribution of the Probability of Guilt.** The median of each distribution is denoted by a black vertical line, which also corresponds to the fixed values used for the Bayesian analysis summarized in Table 1.



**Figure 2. Posterior Distribution of the Probability of Guilt for a Bayesian Analysis Performed Using the Distributions Depicted in Figure 1**



The resultant probability of guilt is found by the mathematical operation of integrating the previous equations with respect to these distributions. It is convenient to perform this integration numerically, and we do so with a technique called “Monte Carlo simulation.”<sup>90</sup> Here is how it works. On each iteration a value is sampled from each of the distributions in Figure 1. Then with these values, the equations shown in Table 1 are used to compute the posterior odds of guilt for that iteration. This process is continued for 100,000 iterations. Because the set of inputs differs for each iteration, a different estimate of the posterior probability of guilt was produced each time, and the total resultant is a posterior distribution of the probability of guilt. Figure 2 shows this posterior, and it summarizes the results of this analysis.

Note that despite considerable variability in the inputs (reflecting uncertainty as to their true values), the posterior probability of guilt ranges from  $\sim 0$  to  $\sim .20$  (i.e., estimated probability of innocence ranges from  $\sim 1$  to  $\sim .80$ ). Both the mean and the mode of this posterior distribution come to  $\sim .03$ . Thus, even after allowing for considerable uncertainty in the inputs to the analysis, the take-home message remains the same: the DNA evidence in this case point strongly in the direction of innocence.

## IX. OTHER EVIDENCE

The preceding analysis considered only DNA evidence because our reading of the case suggests that the DNA match was decisive. For example, the prosecutor’s extensive closing arguments were almost entirely focused on the DNA evidence.<sup>91</sup> Moreover, only DNA evidence was considered because it illustrates how even a small rate of contamination can reverse the conclusion of a DNA match. However, once concerns about the DNA analysis are appreciated, it seems natural to wonder about other evidence—and other possible suspects—related to this case. We highlight here how additional considerations may be incorporated into our analyses.

Before the DNA match in the Michigan State Police Laboratory implicated Gary Leiterman, Jane Mixer was widely believed to have been murdered by John Norman Collins.<sup>92</sup> He was, after all, a serial killer known to have murdered several young women in the Ann Arbor area beginning in the summer of 1967 and ending in the summer of 1969 (Mixer was murdered in March 1969).<sup>93</sup> Figure 3 shows where the victims of the 7 murders originally attributed to Collins were found, and there is nothing in that map to suggest that Mixer was killed by someone else. Even though the prior probability of Collins being the murderer is clearly higher than that of the other adult males who lived in the area in 1969,

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90. See generally Christopher Z. Mooney, *Monte Carlo Simulation* (Sage Univ. Paper Series on Quantitative Applications in the Soc. Sciences, Series No. 07-116, 1997).

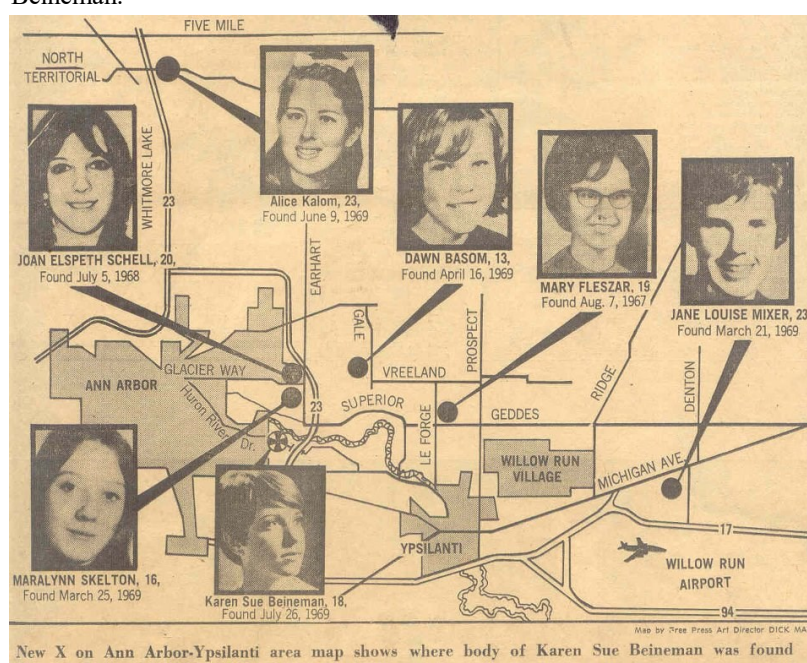
91. Transcript of Record at 14–22, *supra* note 71 (closing statement of Steven Hiller, Washtenaw County Deputy Chief Assistant Prosecuting Attorney).

92. See JAMES, *supra* note 71.

93. See Kim Kozlowski, ‘Michigan Murders’ of Late ‘60s Get Second Look, DETROIT NEWS (Sept. 27, 2016, 12:01 AM), <http://www.detroitnews.com/story/news/michigan/2016/09/27/michigan-murders-book/91151924/> [https://perma.cc/AS4Y-VTVT]; see generally FOURNIER, *supra* note 13.

that fact was not taken into consideration in the analysis performed to this point. Thus far, we have assumed that for any randomly selected adult male living in the vicinity of the murder in 1969, including Collins, the prior probability that he was at the crime scene (and, by assumption, murdered Jane Mixer) was 1-in-1 million, or  $10^{-6}$ .

**Figure 3. A Map Showing Where the Suspected Victims of John Norman Collins (and Jane Mixer) Were Found.**<sup>94</sup> Alice Kalom and Jane Mixer were both University of Michigan graduate students, and both were shot in the head with a .22. Mixer was the 3rd of the 7 murder victims shown on this map. John Norman Collins was convicted of murdering the last of the 7, Karen Sue Beineman.



New X on Ann Arbor-Ypsilanti area map shows where body of Karen Sue Beineman was found  
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In truth, the prior probability was not evenly distributed over the relevant population because Collins was widely believed to have murdered Jane Mixer before the DNA test results became known.<sup>95</sup> Therefore, the prior probability that Collins was at the crime scene was considerably higher than  $10^{-6}$ . That being the case, the prior probability associated with everyone else in the relevant

94. The map was originally published in Walker Lundy & Tom DeLisle, *EMU Coed Found Strangled*, DETROIT FREE PRESS, July 28, 1969, at 4-B.

95. See Kozlowski, *supra* note 93.



population, including Leiterman, was considerably lower than  $10^{-6}$ . Let  $p$  represent the prior probability that Collins was at the crime scene in 1969. The prior probability for all remaining  $(10^6 - 1)$  males would now become  $(1 - p) \times 10^{-6} / (1 - 10^{-6})$ . However, if  $p = .50$ , then the prior probability for everyone else, including Leiterman, becomes  $(1 - .50) \times 10^{-6} / (1 - 10^{-6}) \approx .50 \times 10^{-6}$  (about 1-in-2 million rather than 1-in-1 million). In other words,  $p(H_1)$  in Table 1 would change from its current setting of  $10^{-6}$  to approximately  $.50 \times 10^{-6}$ .

With the prior probability for Collins set to  $p = .50$  and the prior probabilities for the remaining males in the relevant population adjusted to  $(1 - p) \times 10^{-6} / (1 - 10^{-6})$ , the posterior odds and the posterior probability that Leiterman deposited his DNA at the crime scene in 1969 can be re-computed following the steps shown in Appendix B. The posterior odds would change from .0291 (the value shown in Table 1 and Appendix B) to .0146, and the posterior *probability* would equal  $P(\text{Guilt}) = .0144$ , about half of the value we estimated previously (.0283 in Table 1 and Appendix B).

We know of no empirical evidence that could be brought to bear on the estimate of  $p$ , which is why we performed the analysis with  $p$  set to  $10^{-6}$ . However, setting it to a value like .50 does not seem entirely unreasonable. It would mean that, if it happened to be available, the relevant empirical evidence would show that when a murder victim who fits the profile of the other victims of a serial killer is found in the middle of that serial killer's murder spree and in the same general vicinity as the other victims of that serial killer, then, about half the time, that victim, too, will turn out to have been murdered by the serial killer.<sup>96</sup> Whatever the actual value of  $p$  might be, it seems certain that it is much higher than  $10^{-6}$ . By setting it to a more realistic value, the posterior estimate of  $P(\text{Guilt})$  can be updated, as it was above. Additionally, we can also obtain some idea of the relative posterior probability of Collins versus Leiterman being the one who murdered Jane Mixer.

Computing the relative posterior probability for Collins versus Leiterman requires not only that the posterior probability for Leiterman be computed, but also the posterior probability for Collins (i.e., the probability that Collins was at the crime scene despite the DNA match to Leiterman). Note that for any setting of  $p$  equal to or greater than  $10^{-6}$ , the posterior probability of guilt for Leiterman, while always small, will still be much larger than his prior probability. For example, with  $p = 10^{-6}$  for Collins (its original value), the prior probability for Leiterman,  $p(H_1)$ , is also  $10^{-6}$ , whereas the posterior probability for Leiterman (as shown in Table 1) increases to  $P(\text{Guilt}) = .0283$ .

This posterior probability is small, which is the main point of this article, but is also much larger than the prior probability of  $10^{-6}$ . Thus, while the probability that he left his DNA at the crime scene is slight, it is clearly higher than it was before the match. Similarly, as described above, if  $p = .5$  (an arguably more reasonable prior for Collins), the prior probability for Leiterman,  $p(H_1)$ , becomes  $(1 - .5) \times 10^{-6} / (1 - 10^{-6}) \approx .5 \times 10^{-6}$ , and the posterior probability,

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96. See *supra* Figure 3.

$P(\text{Guilt})$ , becomes .0144. Again, despite being small, this posterior probability is still much larger than the prior probability. Because the posterior probability for Leiterman is much larger than his prior probability for any reasonable setting of  $p$ , the posterior probability for the rest of the relevant population, including the posterior probability for Collins, would now have to decrease slightly so that the posterior probabilities over the entire relevant population would sum to 1.

The posterior probability for everyone other than Leiterman is obtained by multiplying their prior probability by  $1 - [P(\text{Guilt}) - P(H_1)] / [1 - P(H_1)]$ , where both  $P(\text{Guilt})$  and  $P(H_1)$  are values that correspond to Leiterman and that depend on  $p$ . Thus, if we again set  $p = .5$  such that  $P(\text{Guilt}) = .0144$  and  $P(H_1) \approx .50 \times 10^{-6}$ , this equation indicates that the prior probabilities for everyone else would be multiplied by  $1 - (.0144 - .50 \times 10^{-6}) / (1 - .50 \times 10^{-6}) = .9856$ . For Collins, his prior probability of  $p = .50$  would drop to a posterior probability of  $.50 \times .9856 = .4928$  as a result of the Leiterman DNA match.

We can now use these values to compute the posterior odds that Collins versus Leiterman was at the crime scene (and, by assumption, murdered Jane Mixer) in 1969. With  $p = .50$ , the posterior odds that Collins versus Leiterman was at the crime scene in 1969 would be  $.4928 / .0144 = 34.22$ . In other words, Collins was ~34 times more likely than Leiterman to have been at the crime scene, despite the DNA match to Leiterman. With  $p = .10$  (which seems like an implausibly low setting given that Collins alone was strongly suspected of having murdered Jane Mixer before the DNA test results became known), the posterior odds that Collins versus Leiterman was at the crime scene in 1969 would be  $.0974 / .0255 = 3.82$  (i.e., Collins would still be ~4 times more likely than Leiterman to have been at the crime scene). The value of  $p$  would have to be set as low as .0285 for the odds to become even.

No matter what the estimate of  $p$ , one advantage of explicitly considering the role of Collins is that it allows us to address the implications of what facially appears to be a somewhat incriminating piece of non-DNA evidence against Leiterman—namely, the fact that in 1969, he owned a .22 caliber gun, which was the kind of gun that was used to kill Jane Mixer.<sup>97</sup> However, John Norman Collins also owned a .22 caliber gun at the time, one that he used to kill another University of Michigan graduate student (by shooting her in the head) just 10 weeks after someone killed University of Michigan graduate student Jane Mixer (also by shooting her in the head with a .22).<sup>98</sup> Thus, ownership of a .22 caliber gun in 1969, while constituting some evidence against both Leiterman and Collins, does not tip the scales in the direction of either one as the person who murdered Jane Mixer. Instead, this evidence variable simply cancels out.

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97. See Transcript of Record, *supra* note 71, at 15.

98. *Id.*



The bottom line of our statistical inquiry into this matter is that if we assume a contamination rate in the Michigan State Police Laboratory in 2002 of 0—despite the disconcerting match to John Ruelas—then the outcome of the analysis suggests that Gary Leiterman deposited his DNA on the Mixer evidence in 1969 (which in turn would mean that he very likely murdered her). However, assuming a contamination rate as low as 1-in-1500 (99.93% of analyses are error-free), then the outcome of the analysis suggests that Gary Leiterman's DNA was deposited on the Mixer evidence through contamination in 2002 with probability .97 (which in turn would mean that John Norman Collins very likely murdered her). The complete reversal of the implication of a contamination rate of 0 and a very small contamination rate is utterly nonintuitive and may help to explain why Leiterman was convicted by a jury.

The analysis summarized in Table 1 is dependent on our reading of the empirical evidence used to estimate the various parameters upon which the analysis was based. If our reading of that evidence turns out to be incorrect, or if new findings come to light showing that the evidence we relied upon is not valid, then the result of the analysis summarized in Table 1 and Figure 2 would change accordingly.

In addition, our analysis is dependent on the assumption that the pertinent events are the four events listed in Table 1. If other events are deemed to be relevant, then they would need to be taken into consideration as well, potentially changing the outcome of the analysis. However, these considerations do not change the main point of the analysis. Under the right conditions, the difference between no chance of contamination and a small chance of contamination can be the difference between no chance of innocence and almost certain innocence.

It is important to emphasize that this analysis does not in any way call DNA evidence into doubt as a general rule. The circumstances of the Leiterman case were special, though perhaps not unique. Had the two DNA analyses been performed in different labs, or at very different times, as would often be the case, the contamination rate could well be infinitesimal. Likewise, had the DNA from the forensic sample been incompatible with the reference sample (saliva) there would be no concern. Furthermore, if the various details of the relative balance between the victim's DNA on the evidence and the suspect's DNA on the evidence had been different from what it was in Leiterman's case (e.g., if much more of the wearer's than the toucher's DNA had been observed), it would point less to innocence. Finally, if any convincing evidence was discovered that made someone a compelling suspect after the DNA match occurred, then the case could be sufficient for a guilty verdict, even with some nonzero chance of an initial contamination. But in the Leiterman case, the relevant data aligns and clearly points in the direction of contamination and therefore, in the direction of innocence.

## APPENDIX A

An alternative approach to estimating the value of  $N$  can be obtained use DNA database statistics. This estimate of  $N$  will represent the size of the active criminal population in Michigan in 1969, a subset of which had their DNA profiles in the database when Leiterman's profile matched in 2004. The Michigan DNA database today, which is part of CODIS, currently contains 377,001 known offender profiles plus 65,971 known arrestee profiles for a total of  $377,001 + 65,971 = 442,972$  known profiles.<sup>99</sup> When an unknown profile from crime scene evidence is entered into this database, and a match is obtained, an investigation is aided.<sup>100</sup> When no match occurs, the still unknown profile is added to the CODIS database as a nonidentified "forensic profile."<sup>101</sup> Currently, in Michigan, 13,132 investigations have been aided from matches to known profiles in the database.<sup>102</sup> In addition, there are 27,729 unknown forensic profiles in the database.<sup>103</sup> Thus, of the  $13,132 + 27,729 = 40,861$  searches,  $13,132 / 40,861 = .32$  matched a known profile (the kind of match that happened when the unknown profile from the Mixer evidence matched the known profile in the database of Gary Leiterman).

If every member of the active criminal population in Michigan today were already in the database, then an unknown forensic profile, when recovered from a crime scene and entered into the database, would always match the known profile of the perpetrator. In that case, it would follow that  $N$  (the size of the active criminal population in Michigan before knowing who matched) would equal the current number of profiles in the database, or  $N = 442,972$ . However, a match occurs only about 1/3 of the time. This result suggests that a better estimate of the active criminal population in Michigan today is about 3 times the number of individuals already in the database, or  $3 \times 442,972 = 1,328,916$  individuals. As noted above, the population of Michigan in 1969 was approximately 89 percent of what it is today, so a reasonable estimate of the active criminal population in Michigan back in 1969 is  $N = .89 \times 1,328,916 = 1,182,735$ .

This estimate of  $N$  uses the highly simplified Lincoln-Peterson mark-recapture formula that has long been used to estimate the size of uncountable wildlife populations, such as the number of fish in a lake.<sup>104</sup> The estimate is only approximate, but it has the virtue of being objective, and the resulting value seems intuitively reasonable. A more exact estimate would require a model that considers the fact that members of active criminal population die, new criminals come of age, different criminals have different probabilities of being captured, and so forth. However, this estimate corresponds closely to the estimate we derived based on geographical considerations ( $N = 1$  million), so it seems like a reasonable figure to use for the main analysis.

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99. *CODIS-NCIS Statistics*, FED. BUREAU INVESTIGATION, <https://www.fbi.gov/services/laboratory/biometric-analysis/codis/ndis-statistics> [<https://perma.cc/NKV4-4KD8>].

100. *Id.*

101. *Id.*

102. *Id.*

103. *Id.*

104. *See generally* BYRON K. WILLIAMS ET AL., ANALYSIS AND MANAGEMENT OF ANIMAL POPULATIONS 290–331 (2002).

## APPENDIX B

A formal calculation of the Bayesian statistics when the third possibility is considered that Letterman both deposited his DNA at the crime scene in 1969 and his DNA contaminated the evidence in the lab in 2002 ( $H_3$ ).

### Hypotheses:

- $H_1$ : DNA deposited at the crime scene in 1969
- $H_2$ : DNA deposited in the lab in 2002
- $H_3$ : Both events occurred

### Bayes' rule:

$$\frac{P(H_i|D)}{P(H_j|D)} = \frac{P(D|H_i)}{P(D|H_j)} \times \frac{(H_i)}{(H_j)},$$

where  $i$  and  $j$  are two hypotheses.

### Events:

Let  $D = E_m \cap E_s \cap E_e \cap E_c$ , where  $A \cap B$  means  $A$  and  $B$ . Then,

$$\frac{P(H_i|D)}{P(H_j|D)} = \frac{P(E_m \cap E_s \cap E_e \cap E_c | H_i)}{P(E_m \cap E_s \cap E_e \cap E_c | H_j)} \times \frac{(H_i)}{(H_j)}.$$

According to the definition of conditional probability,  $P(A \cap B) = P(A|B)P(B)$ . Applying it with  $A = E_s \cap E_e \cap E_c$  and  $B = E_m$ , we obtain  $P(E_m \cap E_s \cap E_e \cap E_c) = P(E_s \cap E_e \cap E_c | E_m)P(E_m)$ . This statement holds conditional on hypotheses as well:

$$P(E_m \cap E_s \cap E_e \cap E_c | H) = P(E_s \cap E_e \cap E_c | E_m, H)P(E_m | H).$$

Conditional on a match ( $m$ ) and a hypothesis ( $H$ ), the probabilities of the forensic DNA being consistent with saliva ( $s$ ), exclusive to Letterman ( $e$ ), and contemporaneously ( $c$ ) analyzed with Letterman's reference DNA are independent. Hence:

$$P(E_m \cap E_s \cap E_e \cap E_c | H) = P(E_s | E_m, H)P(E_e | E_m, H)P(E_c | E_m, H)P(E_m | H).$$

Inserting this expression into Bayes' rule yields:

$$\frac{P(H_i|D)}{P(H_j|D)} = \frac{P(E_s | E_m, H_i)P(E_e | E_m, H_i)P(E_c | E_m, H_i)P(E_m, H_i)}{P(E_s | E_m, H_j)P(E_e | E_m, H_j)P(E_c | E_m, H_j)P(E_m, H_j)} \times \frac{P(H_i)}{P(H_j)} \quad (1)$$

### Priors:

- $P(H_1) = 10^{-6}$
- $P(H_2) = 1/5000 = 2 \times 10^{-4}$
- $P(H_3) = P(H_1) \times P(H_2) = 10^{-6}/5000 = 2 \times 10^{-10}$

**Match:**

- $P(E_m|H_1) = .25$
- $P(E_m|H_2) = 1/1500 = 6.67 \times 10^{-4}$

For the match events under  $H_3$ , there are three different mutually exclusive events of interest: DNA match from the crime scene only ( $E_{m1}$ ), DNA match from the lab only ( $E_{m2}$ ), and DNA match from the crime scene and from the lab ( $E_{m12}$ ). A DNA match from the crime scene would only occur if the DNA was deposited at the crime scene and survived 33 years of storage ( $.5 \times .5 = .25$ ) times the probability that a contamination event did not occur in the lab ( $1499/1500$ ). A DNA match from the lab would only occur if the DNA was not deposited at the crime scene or did not survive 33 years of storage ( $1 - .25 = .75$ ) times the probability that a contamination event did occur in the lab ( $1/1500$ ). And a DNA match from the crime scene *and* from the lab would only occur if the DNA was in fact deposited at the crime scene and survived 33 years of storage ( $.25$ ) times the probability that a contamination event did occur in the lab ( $1/1500$ ). Thus,

- $P(E_{m1}|H_3) = .25(1499/1500) = .2498$
- $P(E_{m2}|H_3) = .75(1/1500) = .0005$
- $P(E_{m12}|H_3) = .25(1/1500) = .00167$

Because these events are exclusive, the total probability that a match event occurs is:

$$\begin{aligned} P(E_m|H_3) &= P(E_{m1}|H_3) + P(E_{m2}|H_3) + P(E_{m12}|H_3) \\ P(E_m|H_3) &= .25(1499/1500) + .75(1/1500) + .25(1/1500) \\ P(E_m|H_3) &= .2505. \end{aligned}$$

**Saliva:**

- $P(E_s|E_m, H_1) = .5$
- $P(E_s|E_m, H_2) = 1$

$P(E_s|E_m, H_1)$  can be more specifically written as  $P(E_s|E_{m1}, H_1)$  to underscore the fact that, under  $H_1$ , the match occurs because the DNA was deposited at the crime scene, not in the lab. Similarly,  $P(E_s|E_m, H_2)$  can be more specifically written as  $P(E_s|E_{m2}, H_2)$  to underscore the fact that, under  $H_2$ , the match occurs because the DNA was deposited in the lab, not at the crime scene. The probability that the DNA was deposited in both places under  $H_3$  is represented as  $P(E_s|E_{m12}, H_3)$ . We note that

- $P(E_s|E_{m1}, H_3) = P(E_s|E_{m1}, H_1) = .5$
- $P(E_s|E_{m2}, H_3) = P(E_s|E_{m2}, H_2) = 1$
- $P(E_s|M_{m12}, H_3) = P(E_s|E_{m1}, H_3) \times P(E_s|E_{m2}, H_3) = .5$

To compute  $P(E_s|M_m, H_3)$  we rely on the conditional form of the law of total probability<sup>105</sup> to specify that

$P(E_s|E_m, H_3) = P(E_s|E_{m1}, E_m, H_3) P(E_{m1}|E_m, H_3)$   
 $(E_s|E_{m1}, E_m, H_3)$  have a redundancy in the conditional and reduce to a simpler form like  $P(E_s|E_{m1}, H_3)$ . In addition, terms like  $P(E_{m1}|E_m, H_3)$  can be written as  $P(E_{m1} \cap E_m|H_3)/P(E_m|H_3) = P(E_{m1}|H_3)/P(E_m|H_3)$ .

Putting it all together yields:

$$\frac{P(E_s|E_m, H_3) = \frac{P(E_s|E_{m1}, H_3) P(E_{m1}|H_3) + P(E_s|E_{m2}, H_3) P(E_{m2}|H_3) + P(E_s|E_{m12}, H_3) P(E_{m12}|H_3)}{P(E_m|H_3)}}{P(E_m|H_3)}$$

The specific values on the right side of this equation were computed above. Substituting those values into this equation yields:

$$P(E_s|M_m, H_3) = .501$$

**Exclusively Leiterman:**

Following the same logic as above:

- $P(E_e|E_m, H_1) = .15$
- $P(E_e|E_m, H_2) = .575$
- $P(E_e|E_{m1}, H_3) = P(E_e|E_m, H_1) = .15$
- $P(E_e|E_{m2}, H_3) = P(E_e|E_m, H_2) = .575$
- $P(E_e|M_{m12}, H_3) = P(E_e|E_{m1}, H_3) \times P(E_e|E_{m2}, H_3) = .08625$
- $P(E_e|M_m, H_3) = .1508$

**Contemporaneous analysis:**

- $P(E_c|E_m, H_1) = 5000/42000 = .119$
- $P(E_c|E_m, H_2) = 1$
- $P(E_c|E_{m1}, H_3) = P(E_c|E_m, H_1) = .119$
- $P(E_c|E_{m2}, H_3) = P(E_c|E_m, H_2) = 1$
- $P(E_c|M_{m12}, H_3) = P(E_c|E_{m1}, H_3) \times P(E_c|E_{m2}, H_3) = .119$
- $P(E_c|M_m, H_3) = .1218$

**Final Calculations:**

Substituting the values computed to this point into Equation 1 to compute the posterior odds of  $H_1$  vs.  $H_2$  yields:

$$\frac{P(H_1|D)}{P(H_2|D)} = 0.0291 \quad 2$$

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105. See SIMON JACKMAN, BAYESIAN ANALYSIS FOR THE SOCIAL SCIENCES 9 (2009).

This is the same value we computed earlier in our main analysis (Table 1). Similarly, substituting the values computed to this point into Equation 1 to compute the posterior odds of  $H_3$  vs.  $H_2$  yields:

$$\frac{P(H_3|D)}{P(H_2|D)} = 5.9645 \times 10^{-6} \quad 3$$

In addition to these values, because the three possibilities are mutually exclusive and exhaustive, it follows that

$$P(H_1|D) + P(H_2|D) + P(H_3|D) = 1 \quad 4$$

Thus, we have three equations (Equations 2 through 4) and 3 unknowns. From the posterior odds of  $H_1$  vs.  $H_2$ , we know that

$$P(H_1|D) = .0291 \times P(H_2|D) \quad 5$$

Similarly, from the posterior odds of  $H_3$  vs.  $H_2$ , we know that

$$P(H_3|D) = 5.9645 \times 10^{-6} \times P(H_2|D) \quad 6$$

Thus, replacing  $P(H_1|D)$  and  $P(H_3|D)$  in Equation 4 with the expressions on the right side of Equations 5 and 6, respectively, we have an equation with one unknown:

$$.0291 \times P(H_2|D) + P(H_2|D) + 5.9645 \times 10^{-6} \times P(H_2|D) = 1$$

Solving for  $P(H_2|D)$  yields:

$$P(H_2|D) = .9717$$

Inserting this value into Equations 5 and 6 above to compute  $P(H_1|D)$  and  $P(H_3|D)$  yields:

$$\begin{aligned} P(H_1|D) &= .0283 \\ P(H_3|D) &= 5.7957 \times 10^{-6} \end{aligned}$$

The probabilities for  $P(H_1|D)$ ,  $P(H_2|D)$  and  $P(H_3|D)$  sum to 1.

Finally, because Leiterman would be guilty if either  $H_1$  or  $H_3$  were true, the probability of guilt is equal to  $P(H_1|D) + P(H_3|D) = .0283 + 5.79 \times 10^{-6} \approx .0283$ , whereas the probability of innocence is equal to  $P(H_2|D) = .9717$ .